

$$\vec{x} = \begin{bmatrix} x^1 \\ x^2 \\ \vdots \\ x^N \end{bmatrix} \quad x^i \text{ is data (trial } i)$$

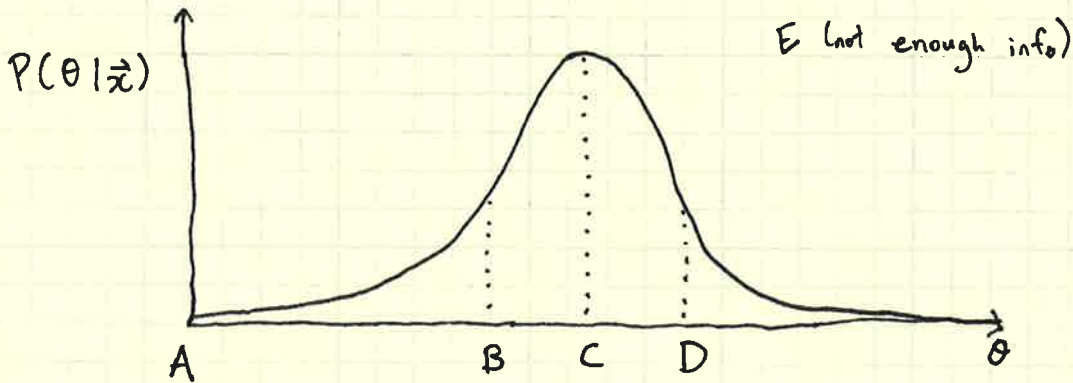
hypothesis:  $P(\vec{x} | \theta)$   $\rightarrow$  model parameter

$$P(\theta | \vec{x}) = \frac{\overbrace{P(\vec{x} | \theta)}^{\text{likelihood}} \overbrace{P(\theta)}^{\text{Prior}}}{\underbrace{P(\vec{x})}_{\text{marginal probability}}}$$

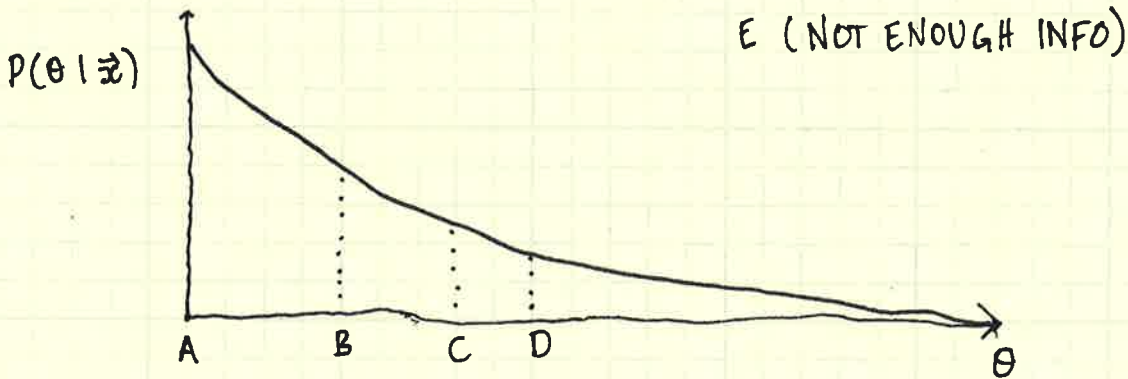
Posterior

Which parameter to choose?

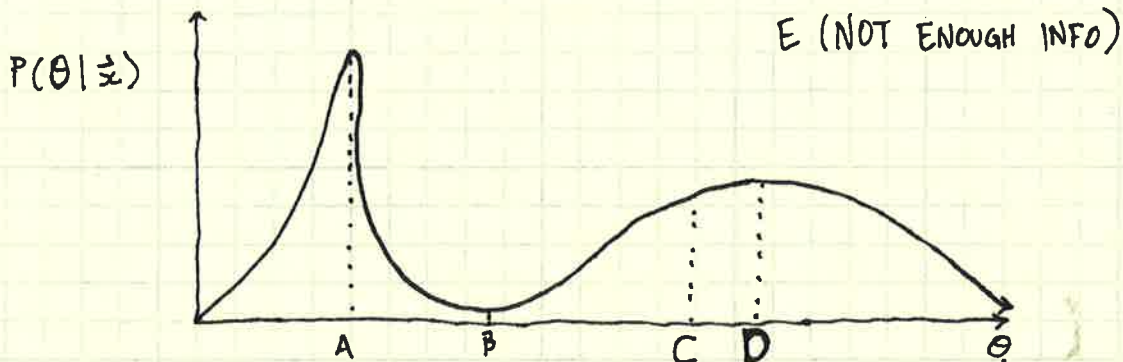
Ex. 1:



Ex 2:



Ex 3:



Maximum a posteriori (MAP) and Maximum Likelihood (ML)

MAXIMIZE  $P(\theta|\vec{x})$ ,  $P(\vec{x}|\theta)$

$$P(\theta|\vec{x}) = \frac{P(\vec{x}|\theta)P(\theta)}{P(\vec{x})}$$

Use MAP:

- 1) Prior is KNOWN and IMPORTANT
- 2) To show  $P(\theta)$  UNIMPORTANT

USE ML:

- 1) Prior does not exist, or UNIMPORTANT
- 2) Convenience, tractability

ML estimate is asymptotically ( $N \rightarrow \infty$ ) [if model is correct]

- UNBIASED ( $\hat{\theta} \rightarrow \theta$ )

- MINIMUM VARIANCE

Example: Fitting a Gaussian

$$P(x|\mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right]$$

Likelihood:  $P(\vec{x}|\mu, \sigma) = \prod_i P(x^i|\mu, \sigma)$  (Independence)

$$= \prod_i \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{(x^i-\mu)^2}{2\sigma^2}\right]$$

$$\log P(\vec{x}|\mu, \sigma) = -\sum_i \left( \log \sigma + \frac{(x^i-\mu)^2}{2\sigma^2} + \frac{1}{2} \log 2\pi \right)$$

{ log monotonic,  
argmax<sub>θ</sub> P(ⓧ|θ) = argmax<sub>θ</sub> log P(ⓧ|θ)

$$\frac{\partial}{\partial \mu} [\log P(\vec{x}|\mu, \sigma)] = \sum_i \frac{x^i - \mu}{\sigma^2} = \frac{1}{\sigma^2} \sum_i x^i - \frac{N}{\sigma^2} \mu$$

= 0 ⇒  
extremum  
condition

$$\hat{\mu} = \frac{1}{N} \sum_i x^i$$

sample mean

$$\frac{\partial}{\partial \sigma} [\log P(\vec{x}|\mu, \sigma)] = -\sum_i \left( \frac{1}{\sigma} - \frac{(x^i-\mu)^2}{\sigma^3} \right)$$

$$= 0 \Rightarrow \hat{\sigma}^2 = \frac{1}{N} \sum_i (x^i - \hat{\mu})^2 \quad \text{sample variance...}$$

$$\langle \hat{\mu} \rangle = \mu \quad (\text{iff data is from the model})$$

$\hat{\mu}$  unbiased

$$\langle \hat{\sigma}^2 \rangle = \left\langle \frac{1}{N} \sum_i (x^i - \hat{\mu})^2 \right\rangle = \left\langle \frac{1}{N} \sum_i ((x^i - \mu) - (\hat{\mu} - \mu))^2 \right\rangle$$

$$= \frac{1}{N} \sum_i \langle (x^i - \mu)^2 \rangle - \frac{2}{N} (\hat{\mu} - \mu) \sum_i \langle (x^i - \mu) \rangle + \frac{1}{N} \langle (\hat{\mu} - \mu)^2 \rangle$$

$$= \frac{1}{N} (N \sigma^2) + \frac{N}{N} \left( \frac{\sigma^2}{N} \right) - \frac{2}{N^2} \sum_{i,j} \langle (x^i - \mu)(x^j - \mu) \rangle$$

Since trials are independent,

$$\langle (x^i - \mu)(x^j - \mu) \rangle = \begin{cases} \sigma^2 & i=j \\ 0 & i \neq j \end{cases}$$

$$\Rightarrow \sigma^2 + \frac{1}{N} \sigma^2 - \frac{2}{N} \sigma^2 = \frac{N-1}{N} \sigma^2$$

BIASED